

EXERCISE 9
KEY

Purpose: To learn more about analyzing multiple regression models (including the LPM) with heteroskedastic errors. **This exercise is due on Tuesday, November 22.**

Work **word problem exercises 1, 2, and 4, page 268** in your textbook. Work **computer exercises C4, C6, and C7 on page 271** in your textbook. Each of these exercises will count 10 points each. I want you to provide “clean” answer sheets in the sense that I want your answers to be “typed up” using Microsoft Word or some equivalent document using a Mac word processor. I will be taking up all of the exercises.

1.

(i) False

(ii) True

(iii) True

2.

$$beer^* = \beta_0 onestar + \beta_1 inc^* + \beta_2 price^* + \beta_3 educ^* + \beta_4 female^* + u^*$$

where $beer^* = beer * \frac{1}{inc}$, $onestar = 1 * \frac{1}{inc}$, $inc^* = \frac{inc}{inc} = 1$, $price^* = price * \frac{1}{inc}$, $educ^* = educ * \frac{1}{inc}$, $female^* = female * \frac{1}{inc}$, $u^* = u * \frac{1}{inc}$.

4.

(i) These variables' coefficients have the expected signs if you believe that, in general, students become more accustomed to college work and, therefore, improve their grades semester by semester.

(ii) Yes, this hypothesis makes sense because the next term's GPA should mirror the previous terms' weighted cumulative GPA. The appropriate t-statistic for this test is $t = (\hat{\beta}_1 - 1)/se(\hat{\beta}_1)$. Using the OLS standard error we have $t = -0.57$. If we use the heteroscedasticity standard error we have $t = -0.60$. Using the two-sigma rule, we accept the null hypothesis that the population coefficient, β_1 , irrespective of which standard error we use.

(iii) The t-statistics are $t = -0.157/0.098 = -1.60$ and $t = -0.157/0.080 = -1.96$, using the OLS and heteroscedasticity-robust standard errors, respectively. The two-tailed probability error of the

second t-statistic is 0.05 while the first t-statistic's two-tailed probability value is greater than 0.05. Yes, the p-values are dependent upon the standard error used.

C4. For the STATA key see Exercise_C4_Ch8_Key.do

The code is

```
* Use VOTE1.dta
regress voteA prtystrA democA lexpndA lexpndB
* Obtain OLS residuals
predict res, residual
* This just shows that the residuals are, by the Least Squares procedure,
* always orthogonal (independent of) the explanatory variables,  $X'uhat = 0$ 
regress res prtystrA democA lexpndA lexpndB
generate res2 = res^2
* The Breusch-Pagan Test for Heteroskedasticity (See equation (8.19), p. 253)
* Use the overall F-statistic from this equation
regress res2 prtystrA democA lexpndA lexpndB
* The special case of White's test (See equation (8.20), p. 253)
regress voteA prtystrA democA lexpndA lexpndB
* Obtain predicted values of  $voteA = voteA\_hat$ 
predict voteA_hat
generate voteA_hat2 = voteA_hat^2
* Use the overall F-statistic from this equation
regress res2 voteA_hat voteA_hat2
```

(i) It is the case that the OLS solution provides the result $X'\hat{u} = 0$. In other words, by construction, the OLS the residuals are orthogonal to the explanatory variables in an OLS regression. That is why a regression of the OLS residuals on the explanatory variables of the regression have an $R^2 = 0$.

(ii) Using Equation (8.19), p. 253, the overall F-statistic of the B-P test equation is 2.33 with a p-value of 0.0581 which is not significant at the 5% level but is at the 10% level. We might say that the incidence of heteroskedasticity in this equation is “marginal.”

(iii) Using Equation (8.20), p. 253, the overall F-statistic of the special case of the White test is 2.79 with p-value of 0.0645. Again, the incidence of heteroscedasticity is “marginal.”

C6. For the STATA key see Exercise_C6_Ch8_Key.do

The code is

```
* Use Crime1.dta
generate arr86 = (narr86>0)
regress arr86 pcnv avgsen tottime ptime86 qemp86
```

```

predict arr86hat
summarize arr86hat
* The weight for each observation using WLS is 1/sqrt[arr86hat*(1 - arr86hat)]
generate weight = 1/[arr86hat*(1 - arr86hat)]^0.5
generate arr86star = arr86*weight
generate onestar = 1*weight
generate pcnvstar = pcnv*weight
generate avgsenstar = avgsen*weight
generate tottimestar = tottime*weight
generate ptime86star = ptime86*weight
generate qemp86star = qemp86*weight
* WLS estimation of the LPM
regress arr86star onestar pcnvstar avgsenstar tottimestar ptime86star qemp86star,
noconstant
test avgsen tottime

```

(i) The smallest fitted value is 0.0066431. The largest fitted value is 0.5576897. Therefore, all of the fitted values are between 0 and 1.

(ii)

```

. * WLS estimation of the LPM
. regress arr86star onestar pcnvstar avgsenstar tottimestar ptime86star qemp86star, noconstant

```

Source	SS	df	MS	Number of obs	=	2,725
Model	1109.44323	6	184.907205	F(6, 2719)	=	188.31
Residual	2669.84459	2,719	.98192151	Prob > F	=	0.0000
				R-squared	=	0.2936
				Adj R-squared	=	0.2920
Total	3779.28782	2,725	1.38689461	Root MSE	=	.99092

arr86star	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
onestar	.4475965	.0179922	24.88	0.000	.4123167 .4828763
pcnvstar	-.1678436	.0189122	-8.87	0.000	-.2049272 -.13076
avgsenstar	.0053665	.0051146	1.05	0.294	-.0046624 .0153954
tottimestar	-.0017615	.0032514	-0.54	0.588	-.008137 .004614
ptime86star	-.0246188	.0030451	-8.08	0.000	-.0305898 -.0186479
qemp86star	-.0451885	.0054225	-8.33	0.000	-.0558212 -.0345558

(iii)

The *avgsen* and *totime* variables are jointly insignificant since the joint F-test statistic has a p-value 0.4129 which is greater than 0.05.

```

. test avgsen tottime

( 1) avgsenstar = 0
( 2) tottimestar = 0

F( 2, 2719) = 0.88
Prob > F = 0.4129

```

C7. For the STATA key see Exercise_C7_Ch8_Key.do

The code is

*** Use Loanapp.dta**

*** LPM with OLS standard errors**

**regress approve white hrat obrat loanprc unem male married dep sch cosign chist pubrec
mortlat1 mortlat2 vr**

*** LPM with heteroskedasticity-robust standard errors**

**regress approve white hrat obrat loanprc unem male married dep sch cosign chist pubrec
mortlat1 mortlat2 vr, vce(robust)**

predict yhat

summarize yhat

(i)

```
. * LPM with OLS standard errors
. regress approve white hrat obrat loanprc unem male married dep sch cosign chist pubrec mortlat1 mortlat2 vr
```

Source	SS	df	MS	Number of obs	=	1,971
Model	35.4004787	15	2.36003192	F(15, 1955)	=	25.86
Residual	178.393534	1,955	.09124989	Prob > F	=	0.0000
				R-squared	=	0.1656
				Adj R-squared	=	0.1592
Total	213.794013	1,970	.10852488	Root MSE	=	.30208

approve	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
white	.1288196	.0197317	6.53	0.000	.0901223 .1675169
hrat	.001833	.0012632	1.45	0.147	-.0006444 .0043104
obrat	-.0054318	.0011018	-4.93	0.000	-.0075926 -.003271
loanprc	-.1473001	.0375159	-3.93	0.000	-.2208755 -.0737246
unem	-.0072989	.003198	-2.28	0.023	-.0135708 -.0010271
male	-.0041441	.0188644	-0.22	0.826	-.0411405 .0328523
married	.0458241	.0163077	2.81	0.005	.0138418 .0778064
dep	-.0068274	.0067013	-1.02	0.308	-.0199699 .0063151
sch	.0017525	.0166498	0.11	0.916	-.0309006 .0344057
cosign	.0097722	.0411394	0.24	0.812	-.0709094 .0904538
chist	.1330267	.0192627	6.91	0.000	.0952492 .1708043
pubrec	-.2419268	.0282274	-8.57	0.000	-.2972858 -.1865677
mortlat1	-.0572511	.050012	-1.14	0.252	-.1553336 .0408314
mortlat2	-.1137234	.0669838	-1.70	0.090	-.2450905 .0176438
vr	-.0314408	.0140313	-2.24	0.025	-.0589586 -.0039229
_cons	.9367312	.0527354	17.76	0.000	.8333077 1.040155

```

. * LPM with heteroskedasticity-robust standard errors
. regress approve white hrat obrat loanprc unem male married dep sch cosign chist pubrec mortlat1 mortlat2 vr, vce(robust)

Linear regression                               Number of obs   =    1,971
                                                F(15, 1955)    =    14.98
                                                Prob > F       =    0.0000
                                                R-squared      =    0.1656
                                                Root MSE      =    .30208

```

approve	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
white	.1288196	.0258693	4.98	0.000	.0780852	.179554
hrat	.001833	.001467	1.25	0.212	-.0010441	.0047101
obrat	-.0054318	.001331	-4.08	0.000	-.0080421	-.0028215
loanprc	-.1473001	.0378351	-3.89	0.000	-.2215013	-.0730988
unem	-.0072989	.0037122	-1.97	0.049	-.0145792	-.0000187
male	-.0041441	.0193044	-0.21	0.830	-.0420035	.0337152
married	.0458241	.0172374	2.66	0.008	.0120186	.0796296
dep	-.0068274	.0069038	-0.99	0.323	-.0203669	.0067122
sch	.0017525	.017146	0.10	0.919	-.0318739	.0353789
cosign	.0097722	.0395825	0.25	0.805	-.0678561	.0874005
chist	.1330267	.0246202	5.40	0.000	.0847421	.1813114
pubrec	-.2419268	.0427922	-5.65	0.000	-.3258498	-.1580037
mortlat1	-.0572511	.0662234	-0.86	0.387	-.1871269	.0726247
mortlat2	-.1137234	.0910697	-1.25	0.212	-.2923274	.0648806
vr	-.0314408	.0144855	-2.17	0.030	-.0598493	-.0030322
_cons	.9367312	.0593886	15.77	0.000	.8202595	1.053203

The robust confidence interval for β_{white} is [0.078, 0.179].

The non-robust confidence interval for β_{white} is [0.090, 0.167].

The non-robust confidence interval is narrower than the robust confidence interval.

(ii)

```
. summarize yhat
```

Variable	Obs	Mean	Std. Dev.	Min	Max
yhat	1,971	.876205	.1340514	.2273447	1.172988

The minimum fitted value is greater than 0 but the maximum fitted value is greater than 1.0 making the application of WLS to this LPM problematic. Either corresponding observations have to be dropped or the fitted values that are outside of the unit interval have to be “trimmed” to be inside the unit interval. For example, if a yhat is less than zero, make it 0.001. If a yhat is greater than one, make it 0.999. After the trimming then apply WLS.