

Lecture 2

①

Basic Statistics Review

Discrete Random Variables versus
Continuous Random Variables.

Discrete RV

Bernoulli (binary) rv

$$\Pr(X=1) = p, \quad \Pr(X=0) = 1-p$$

where $p = \text{probability of success}$ $0 \leq p \leq 1$.

Example: flip of fair coin ($p = \frac{1}{2}$)

Let $X=1$ when head occurs

$X=0$ when tail occurs

Consider the sample space of 3 flips of a fair coin:

$\{H, H, H\}$, $\{H, H, T\}$, $\{H, T, H\}$, $\{T, H, H\}$

$\{H, T, T\}$, $\{T, T, H\}$, $\{T, H, T\}$, $\{T, T, T\}$

These are all possible outcomes in the sample space.

Lecture 2

Let S = no. of successes (heads) in 3 coin tosses.

$$\text{Then } \Pr(S = s = 3) = 1/8$$

$$\Pr(S = s = 2) = 3/8$$

$$\Pr(S = s = 1) = 3/8$$

$$\Pr(S = s = 0) = 1/8$$

Probability Density Function (pdf)

S is a binomial r.v with $N=3$ trials and the probability of success of $p=1/2$.

Then from basic probability theory we know

$$\Pr(S = s) = \binom{N}{s} p^s (1-p)^{N-s}$$

where
$$\binom{N}{s} = \frac{N!}{(N-s)! s!}$$

and $N! = N(N-1)(N-2) \dots 2 \cdot 1$ and similarly for $s!$ and $(N-s)!$ with $0! = 1$.

Lecture 2

In our case $N=3$, $p=\frac{1}{2}$

$$\begin{aligned} \text{Pr}(S=s=3) &= \binom{3}{3} \left(\frac{1}{2}\right)^3 \left(1-\frac{1}{2}\right)^0 \\ &= 1 \left(\frac{1}{2}\right)^3 \left(\frac{1}{2}\right)^0 \\ &= \left(\frac{1}{2}\right)^3 = \frac{1}{8} \end{aligned}$$

$$\begin{aligned} \text{Pr}(S=s=2) &= \binom{3}{2} \left(\frac{1}{2}\right)^2 \left(1-\frac{1}{2}\right)^1 \\ &= 3 \left(\frac{1}{2}\right)^3 = \frac{3}{8} \end{aligned}$$

$$\begin{aligned} \text{Pr}(S=s=1) &= \binom{3}{1} \left(\frac{1}{2}\right)^1 \left(1-\frac{1}{2}\right)^2 \\ &= 3 \left(\frac{1}{2}\right)^3 = \frac{3}{8} \end{aligned}$$

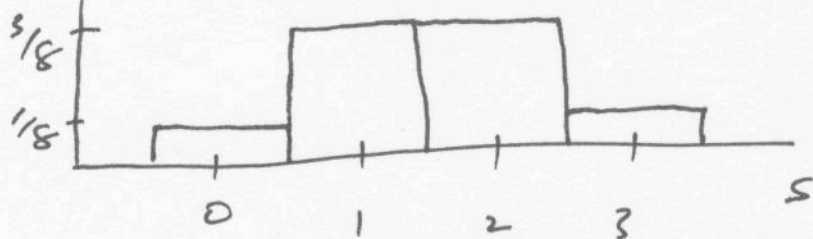
$$\begin{aligned} \text{Pr}(S=s=0) &= \binom{3}{0} \left(\frac{1}{2}\right)^0 \left(1-\frac{1}{2}\right)^3 \\ &= 1 \left(\frac{1}{2}\right)^3 = \frac{1}{8} \end{aligned}$$

Lecture 2

9

A graph of the binomial pdf for $N=3, p=1/2$.

$\Pr(\mathcal{S}=s)$



Notice that the sum of the areas of the rectangles add to one. \mathcal{S} is a discrete rv because \mathcal{S} takes on only a finite (discrete) number of outcomes.

Cumulative Distribution Function (cdf)

$\Pr(\mathcal{S} \leq s)$

Lecture 2

(5)

$$\Pr(S' \leq 0) = \frac{1}{8}$$

$$\begin{aligned}\Pr(S' \leq 1) &= \Pr(S'=0) + \Pr(S'=1) \\ &= \frac{1}{8} + \frac{3}{8} = \frac{1}{2}\end{aligned}$$

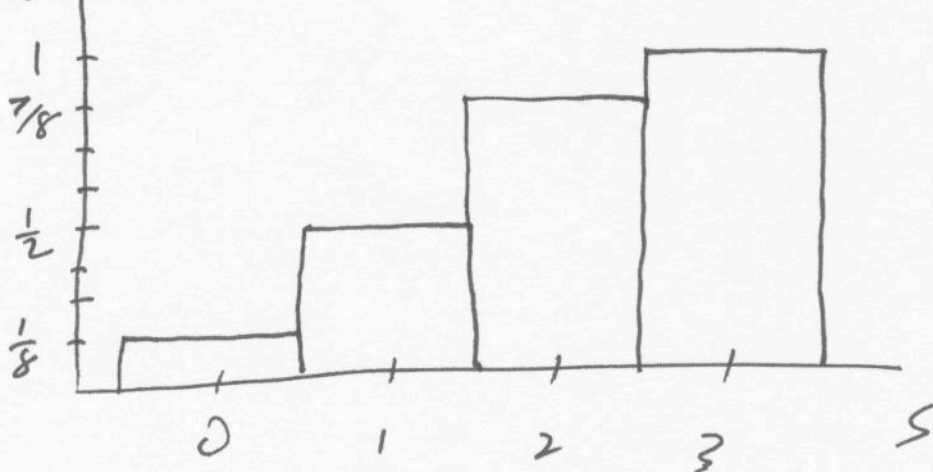
$$\begin{aligned}\Pr(S' \leq 2) &= \Pr(S'=0) + \Pr(S'=1) + \Pr(S'=2) \\ &= \frac{1}{8} + \frac{3}{8} + \frac{3}{8} = \frac{7}{8}\end{aligned}$$

$$\begin{aligned}\Pr(S \leq 3) &= \Pr(S'=0) + \Pr(S'=1) + \Pr(S'=2) \\ &\quad + \Pr(S'=3) = \frac{1}{8} + \frac{3}{8} + \frac{3}{8} + \frac{1}{8} = 1\end{aligned}$$

Here is a graph of the Cumulative Distribution

Function:

$\Pr(S' \leq s)$



Here the areas don't sum to one but the rectangles eventually grow to have an area of one.

Lecture 2

Let us turn to the derivations of the mean and variance of a discrete rv.

For the discrete rv X with possible outcomes x_1, x_2, \dots, x_r

The expected value of X , $E(X)$ is calculated as

$$E(X) = \sum_{i=1}^r x_i \cdot \Pr(X=x_i)$$

The variance of X is calculated as

$$V(X) = \sum_{i=1}^r (x_i - E(X))^2 \cdot \Pr(X=x_i)$$

For the three coin toss problem

$$E(S^1) = \sum_{s=0}^3 s \cdot \Pr(S^1=s)$$

$$= 0 \cdot \Pr(S^1=0) + 1 \cdot \Pr(S^1=1) + 2 \cdot \Pr(S^1=2) + 3 \cdot \Pr(S^1=3)$$

$$= 0 \cdot \frac{1}{8} + 1 \cdot \frac{3}{8} + 2 \cdot \frac{3}{8} + 3 \cdot \frac{1}{8}$$

$$= \frac{3+6+3}{8} = 1.5$$

(7)

Lecture 2

$$\begin{aligned}
 V(S') &= \sum_{s=0}^3 (s-1.5)^2 \cdot \Pr(S'=s) \\
 &= (0-1.5)^2 \cdot \Pr(S'=0) + (1-1.5)^2 \cdot \Pr(S'=1) \\
 &\quad + (2-1.5)^2 \cdot \Pr(S'=2) + (3-1.5)^2 \cdot \Pr(S'=3) \\
 &= (-1.5)^2 \cdot \frac{1}{8} + (-0.5)^2 \cdot \frac{3}{8} + (0.5)^2 \cdot \frac{3}{8} + (1.5)^2 \cdot \frac{1}{8} \\
 &= 0.75
 \end{aligned}$$

Continuous Probability Density Function (pdf)

The continuous pdfs are different than the discrete pdfs. In the continuous pdf the probability of the continuous rv taking on the specific value of a point is zero, i.e.

$\Pr(X=x_0) = 0$. Only intervals like

$(x_1 < X < x_2)$ have probability. Let $f(x)$

be the pdf of a continuous rv X . Then

$$(i) \int_{-\infty}^{\infty} f(x) dx = 1$$

(f)

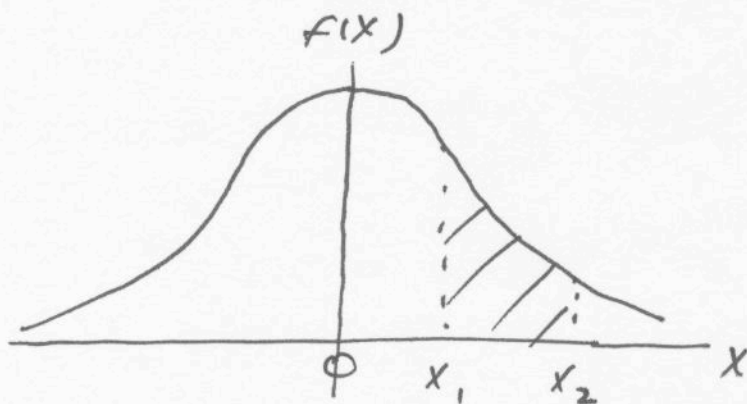
Lecture 2

$$(ii) \Pr(x_1 < X < x_2) = \int_{x_1}^{x_2} f(x) dx$$

Example: $X \sim N(0, 1)$

Then

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$$



$$\Pr(x_1 < X < x_2) = \int_{x_1}^{x_2} f(x) dx = \int_{x_1}^{x_2} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

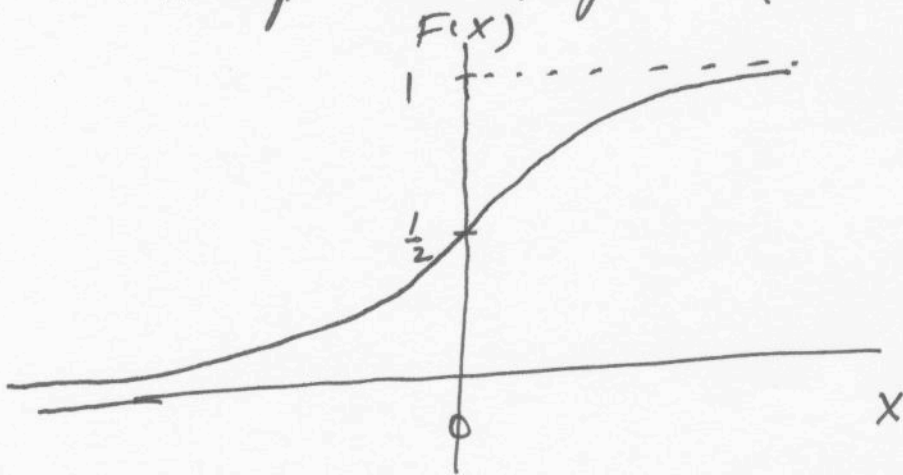
= area under the pdf from x_1 to x_2 .

The cumulative distribution function for a continuous r.v. X is defined as

Lecture 2

$$F(x) = \Pr(X \leq x) = \int_{-\infty}^x f(x) dx$$

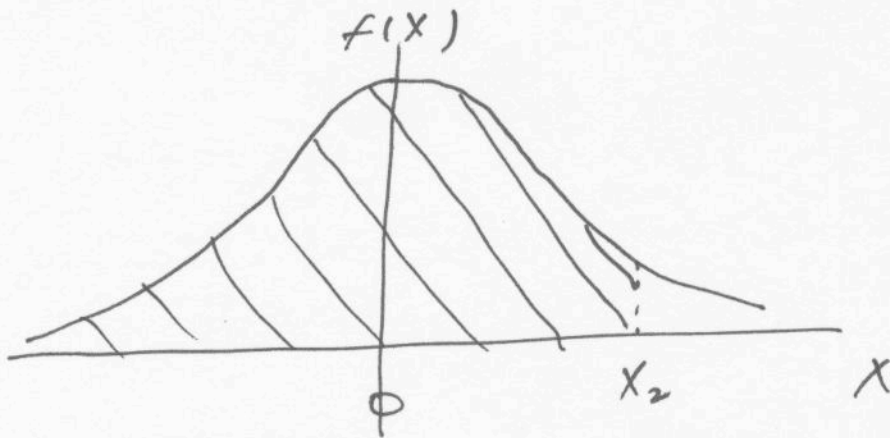
and is represented graphically by (for the $X \sim N(0,1)$)



If $X \sim N(0,1)$ you can calculate

$$\Pr(x_1 < X < x_2) = F(x_2) - F(x_1)$$

where $F(x_2)$ looks like



the shaded area under the pdf.

Lecture 2

(10)

Example (use your standard normal table in the back of your Wooldridge book). Let $X \sim N(0, 1)$

$$\begin{aligned}\Pr(1 < X < 2) &= F(2) - F(1) \\ &= 0.9772 - 0.8413 \\ &= 0.1359\end{aligned}$$

In a similar vein,

$$\text{Let } X \sim N(50, 15)$$

Calculate

$$\Pr(45 < X < 55)$$

$$\text{Let } z = \frac{X - \mu}{\sigma}$$

$$\text{Then } \Pr(45 < X < 55) = \Pr\left(\frac{45 - 50}{\sqrt{15}} < \frac{X - 50}{\sqrt{15}} < \frac{55 - 50}{\sqrt{15}}\right)$$

$$= \Pr\left(\frac{-5}{\sqrt{15}} < z < \frac{5}{\sqrt{15}}\right)$$

$$= \Pr\left(\frac{-5}{3.873} < z < \frac{5}{3.873}\right) = \Pr(1.29 < z < 1.29)$$

$$= \frac{1}{2}(1.29) - \frac{1}{2}(-1.29) = 0.9015 - 0.0985$$

$$= 0.803$$

Lecture 2

(11)

For a continuous rv X the population mean is calculated as

$$E(X) = \int_{-\infty}^{\infty} x \cdot f(x) dx$$

The variance is calculated as

$$V(X) = \int_{-\infty}^{\infty} (x - E(X))^2 f(x) dx.$$

Some popular continuous rv are:

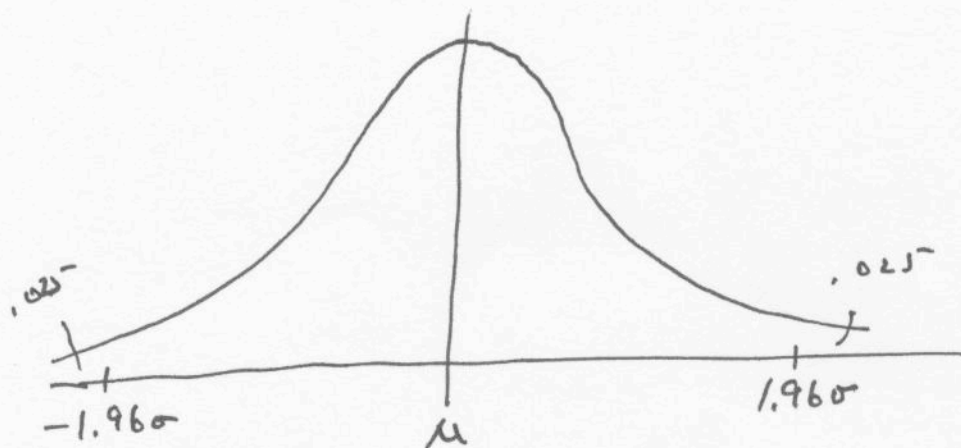
1) $X \sim N(\mu, \sigma^2)$

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}}$$

and looks like

Lecture 2

(12)



see the next page for
a summary of some
popular continuous
pdfs.

Continuous Distributions

Distribution	pdf	mean	Variance
Uniform $X \sim U(a, b)$	$f(x) = \frac{1}{b-a}$ for $(a < x < b)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Normal $X \sim N(\mu, \sigma^2)$	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}}$ for $(-\infty < x < \infty)$	μ	σ^2
Exponential	$f(x) = \frac{1}{\beta} e^{-x/\beta}$ $\beta > 0, 0 \leq x < \infty$	β	β^2
Gamma	$f(x) = \left[\frac{1}{\Gamma(\alpha)\beta^\alpha} \right] x^{\alpha-1} e^{-x/\beta}$	$\alpha\beta$	$\alpha\beta^2$
Chi-square	$f(x^2) = \frac{(x^2)^{(v/2)-1} e^{-x^2/2}}{2^{v/2} \Gamma(v/2)}$	v	$2v$
Beta	$f(x) = \left[\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \right] x^{\alpha-1} (1-x)^{\beta-1}$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$
t-distribution	$T = \frac{Z}{\sqrt{\chi^2/v}}$ and as $v \rightarrow \infty T \rightarrow Z$ r.v.		

Lecture 2

(14)

Continuous Distributions (continued)

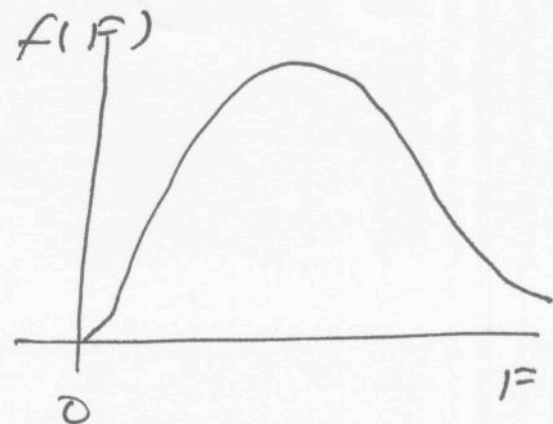
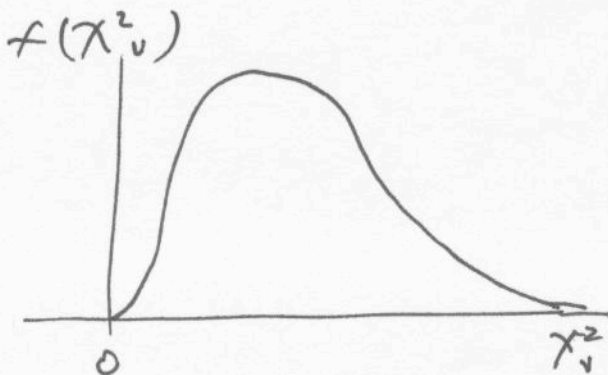
Distribution	pdf	mean	variance
--------------	-----	------	----------

t-distribution	$f(x) = \frac{\Gamma[(\nu+1)/2]}{\Gamma(\nu/2)} \cdot \frac{1}{\sqrt{\nu\pi}} \cdot \frac{1}{(1+x^2/\nu)^{(\nu+1)/2}}$ <p>$-\infty < x < \infty$</p>		$\frac{\nu}{\nu-2}$ $\frac{\nu}{\nu-2}$
----------------	---	--	---

The t-distribution looks like the $N(0,1)$ distribution but has "fatter" tails.

F-distribution
 m = numerator df
 n = denominator df

$$f(x) = \text{Beta density with } \alpha = \frac{m}{2} \quad \beta = \frac{n}{2}$$



χ^2 and F look similar.

Lecture 2

(15)

Covariance : $E(X - \mu_X)(Y - \mu_Y)$

Correlation : $\rho_{XY} = \frac{E(X - \mu_X)(Y - \mu_Y)}{\sqrt{\text{Var}(X) \cdot \text{Var}(Y)}}$

$$-1 \leq \rho_{XY} \leq 1$$

Properties of Expectations (both discrete and continuous pdfs)

(i) $E(c) = c$, c is constant

(ii) $E(ax + by + c) = aE(X) + bE(Y) + c$

(iii) $E\sum_{i=1}^N a_i X_i = \sum_{i=1}^N a_i E(X_i)$

Properties of Variance

(i) $\text{Var}(ax + by + c) = a^2 \text{Var}(X) + b^2 \text{Var}(Y) + 2ab \text{Cov}(X, Y)$

(ii) $\text{Var}(c) = 0$

Properties of Covariance

$$(i) \text{cov}(a_1x + b_1, a_2y + b_2) \\ = a_1a_2 \text{cov}(x, y)$$

(ii) When x and y are independent

$$\text{cov}(x, y) = 0.$$